

Modeling and Verification of Probabilistic Systems

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<http://moves.rwth-aachen.de/teaching/ws-1516/movep15/>

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Overview

- 1 Recall: continuous-time Markov chains
- 2 Transient distribution
- 3 Uniformization
- 4 Strong and weak bisimulation
- 5 Computing transient probabilities
- 6 Summary

Negative exponential distribution

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Density of exponential distribution

The density of an *exponentially distributed* r.v. Y with *rate* $\lambda \in \mathbb{R}_{>0}$ is:

$$f_Y(x) = \lambda \cdot e^{-\lambda \cdot x} \quad \text{for } x > 0 \quad \text{and } f_Y(x) = 0 \text{ otherwise}$$

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► Expectation $E[Y] = \int_0^\infty x \cdot \lambda \cdot e^{-\lambda \cdot x} dx = \frac{1}{\lambda}$

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- ▶ Expectation $E[Y] = \int_0^\infty x \cdot \lambda \cdot e^{-\lambda \cdot x} dx = \frac{1}{\lambda}$
- ▶ Variance $Var[Y] = \int_0^\infty (x - E[X])^2 \lambda \cdot e^{-\lambda \cdot x} dx = \frac{1}{\lambda^2}$

Continuous-time Markov chain

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A **CTMC** is a tuple $(S, \mathbf{P}, r, \iota_{\text{init}}, AP, L)$ where

- ▶ $(S, \mathbf{P}, \iota_{\text{init}}, AP, L)$ is a DTMC, and
- ▶ $r : S \rightarrow \mathbb{R}_{>0}$, the **exit-rate function**

Let $\mathbf{R}(s, s') = \mathbf{P}(s, s') \cdot r(s)$ be the transition rate of transition (s, s')

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- ▶ **residence** time in state s is exponentially distributed with **rate** $r(s)$.
- ▶ phrased alternatively, the **average** residence time of state s is $\frac{1}{r(s)}$.

CTMC semantics

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Residence time distribution

The probability to *take some* outgoing transition from s in $[0, t]$ is:

$$\int_0^t r(s) \cdot e^{-r(s) \cdot x} dx = 1 - e^{-r(s) \cdot t}$$

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Theorem: transient distribution as linear differential equation

The **transient** probability vector $\underline{p}(t) = (p_{s_1}(t), \dots, p_{s_k}(t))$ satisfies:

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$$\underline{p}'(t) = \underline{p}(t) \cdot (\mathbf{R} - \mathbf{r}) \quad \text{given} \quad \underline{p}(0)$$

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Proof:

On the blackboard.

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But: [numerical instability](#) due to fill-in of $(\mathbf{R}-\mathbf{r})^i$ in presence of positive and negative entries in the matrix $\mathbf{R}-\mathbf{r}$.

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It follows that $\bar{\mathbf{P}}$ is a stochastic matrix and $\text{unif}(r, \mathcal{C})$ is a CTMC.

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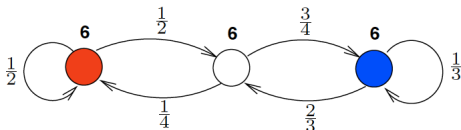
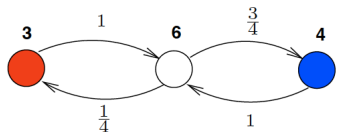
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CTMC \mathcal{C} and its uniformized counterpart $\text{unif}(6, \mathcal{C})$

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- ▶ Then **normalize** the residence time of all states with respect to r as follows:
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 2. decrease the transition probabilities by a factor $\frac{r(s)}{r}$, and

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$$\bar{\mathbf{P}}(s, s') = \frac{r(s)}{r} \cdot \mathbf{P}(s, s') \text{ if } s' \neq s \quad \text{and} \quad \bar{\mathbf{P}}(s, s) = \frac{r(s)}{r} \cdot \mathbf{P}(s, s) + 1 - \frac{r(s)}{r}.$$

Intuition

- ▶ Fix all exit rates to (at least) the **maximal** exit rate r occurring in CTMC \mathcal{C} .
- ▶ Thus, $\frac{1}{r}$ is the **shortest** mean residence time in the CTMC \mathcal{C} .
- ▶ Then **normalize** the residence time of all states with respect to r as follows:
 1. replace an average residence time $\frac{1}{r(s)}$ by a shorter (or equal) one, $\frac{1}{r}$
 2. decrease the transition probabilities by a factor $\frac{r(s)}{r}$, and
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Uniformization: intuition

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That is, **slow down** state s whenever $r(s) < r$.

Overview

- 1 Recall: continuous-time Markov chains
- 2 Transient distribution
- 3 Uniformization
- 4 Strong and weak bisimulation**
- 5 Computing transient probabilities
- 6 Summary

Strong bisimulation on DTMCs

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Probabilistic bisimulation

[Larsen & Skou, 1989]

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The last two conditions amount to $\mathbf{R}(s, C) = \mathbf{R}(t, C)$ for all equivalence classes $C \in S/R$.

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For states in R , the **conditional** probability of moving by a single transition to **another** equivalence class is equal. In addition, either all states in an equivalence class C almost surely stay there, or have an option to escape from C .

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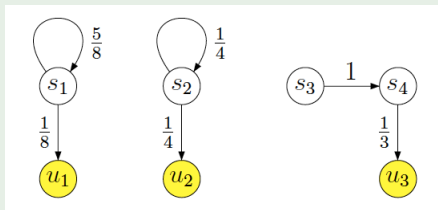
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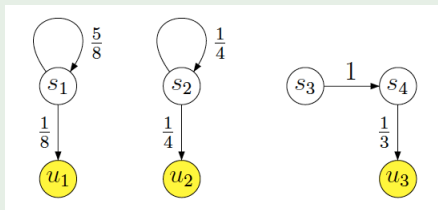
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Weak bisimulation on DTMC: example

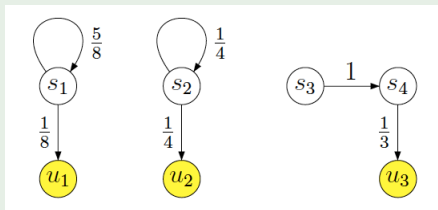


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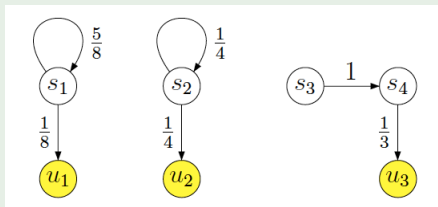
The equivalence relation R with $S/R = \{ \{s_1, s_2, s_3, s_4\}, \{u_1, u_2, u_3\} \}$ is a weak bisimulation.

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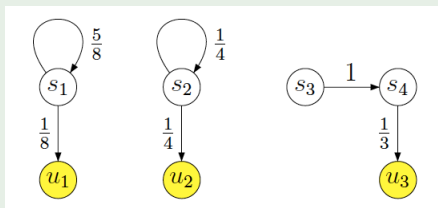
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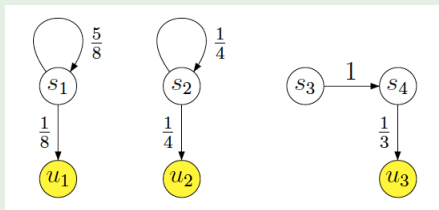
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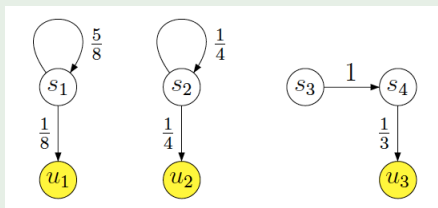
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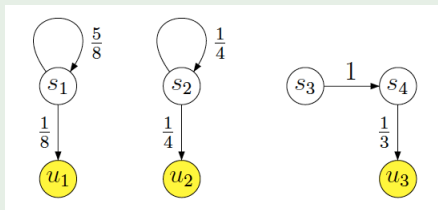
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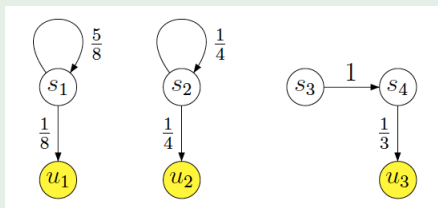
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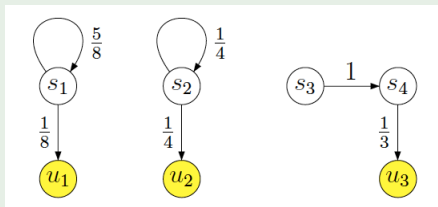


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$$\frac{\mathbf{P}(s_1, C)}{1 - \mathbf{P}(s_1, [s_1]_R)} = \frac{1/8}{1 - 5/8} = \frac{1/4}{1 - 1/4} = \frac{\mathbf{P}(s_2, C)}{1 - \mathbf{P}(s_2, [s_2]_R)} = \frac{1/3}{1} = \frac{\mathbf{P}(s_4, C)}{1 - \mathbf{P}(s_4, [s_4]_R)}.$$

Note that $\mathbf{P}(s_3, [s_3]_R) = 1$.

Weak bisimulation on DTMC: example



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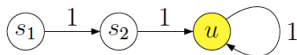
Note that $\mathbf{P}(s_3, [s_3]_R) = 1$. Since s_3 can reach a state outside $[s_3]$ as s_1, s_2 and s_4 , it follows that $s_1 \approx_p s_2 \approx_p s_3 \approx_p s_4$.

Reachability condition

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Remark

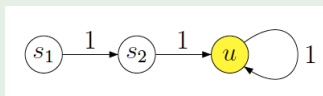
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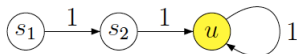


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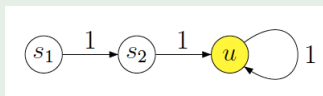


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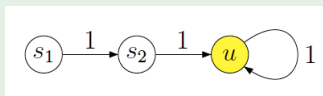


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It is not difficult to establish $s_1 \approx s_2$. Note: $\mathbf{P}(s_1, [s_1]_R) = 1$, but $\mathbf{P}(s_2, [s_2]_R) < 1$. Both s_1 and s_2 can reach a state outside $[s_1]_R = [s_2]_R$. The reachability condition is essential to establish $s_1 \approx s_2$ and cannot be dropped: otherwise s_1 and s_2 would be weakly bisimilar to an equally labelled absorbing state.

Weak bisimulation on CTMCs

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Weak probabilistic bisimulation

[Bravetti, 2002]

Weak bisimulation on CTMCs

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Weak probabilistic bisimilarity

Let \mathcal{C} be a CTMC and s, t states in \mathcal{C} . Then: s is **weak probabilistically bisimilar** to t , denoted $s \approx_m t$, if there **exists** a weak probabilistic bisimulation R with $(s, t) \in R$.

A useful lemma

Let \mathcal{C} be a CTMC and R an equivalence relation on S with $(s, t) \in R$, $\mathbf{P}(s, [s]_R) < 1$ and $\mathbf{P}(t, [t]_R) < 1$.

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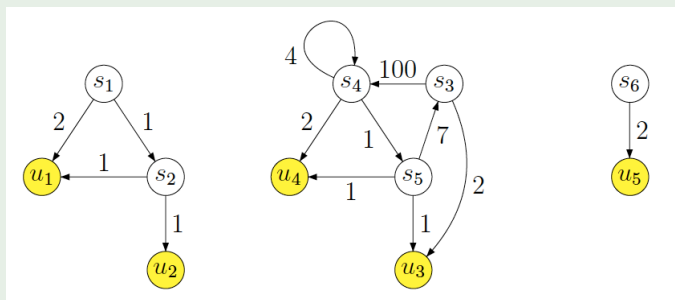
2. $\mathbf{R}(s, C) = \mathbf{R}(t, C)$ for all $C \in S/R$ with $C \neq [s]_R = [t]_R$.

Proof:

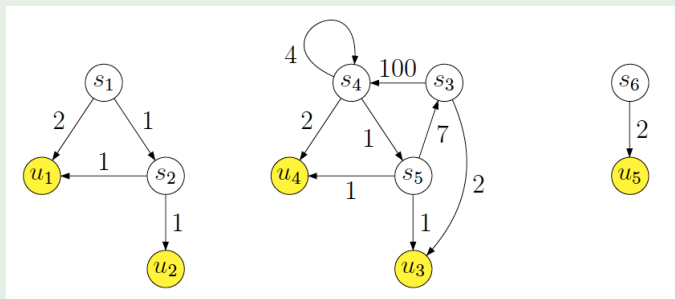
Left as an exercise.

Weak bisimulation on CTMCs: example

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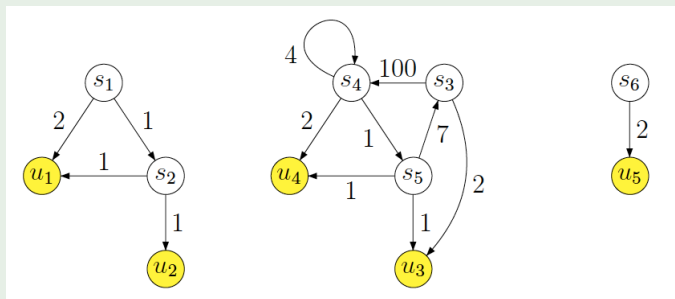


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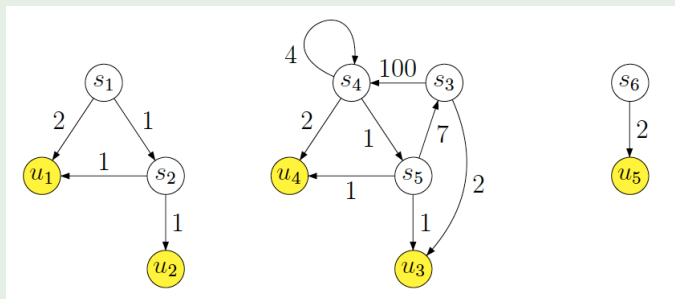
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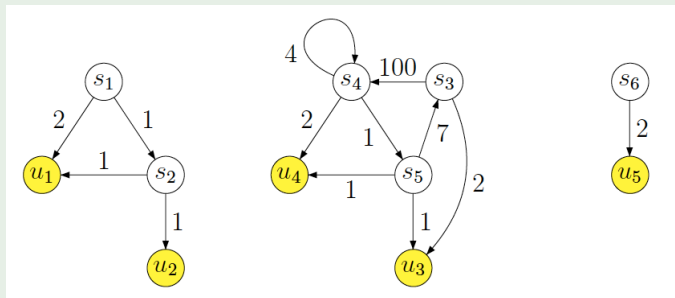
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Properties (without proof)

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Strong and weak bisimulation in uniform CTMCs

For all uniform CTMCs \mathcal{C} and states s, u in \mathcal{C} , we have:

$$s \sim_m u \quad \text{iff} \quad s \approx_m u \quad \text{iff} \quad s \sim_p u.$$

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Preservation of transient probabilities

For all CTMCs \mathcal{C} with states s, u in \mathcal{C} and $t \in \mathbb{R}_{\geq 0}$, we have:

$$s \approx_m u \quad \text{implies} \quad \underline{p}^s(t) = \underline{p}^u(t)$$

where $\underline{p}^s(0) = \mathbf{1}_s$ and $\underline{p}^u(0) = \mathbf{1}_u$ where $\mathbf{1}_s$ is the characteristic function for state s , i.e., $\mathbf{1}_s(s') = 1$ iff $s = s'$.

Overview

- 1 Recall: continuous-time Markov chains
- 2 Transient distribution
- 3 Uniformization
- 4 Strong and weak bisimulation
- 5 Computing transient probabilities**
- 6 Summary

Computing transient probabilities

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The transient probability vector $\underline{p}(t) = (p_{s_1}(t), \dots, p_{s_k}(t))$ satisfies:

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Thus:

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Computing a matrix exponential

Exploit [Taylor-Maclaurin](#) expansion.

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As \bar{P} is a stochastic matrix, computing the matrix exponential \bar{P}^i is numerically stable.

Intermezzo: Poisson distribution

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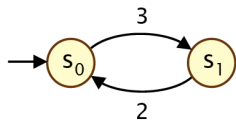
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Remark

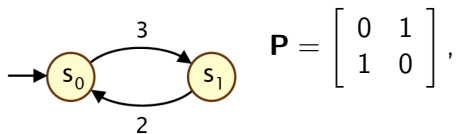
The Poisson distribution can be derived as a limiting case to the binomial distribution as the number of trials goes to infinity and the expected number of successes remains fixed.

Transient probabilities: example

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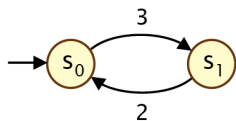


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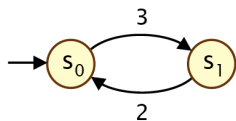
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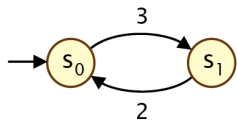
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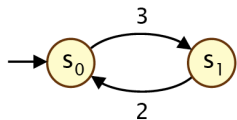
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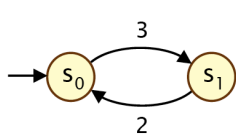


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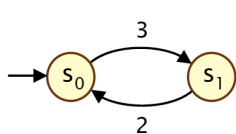


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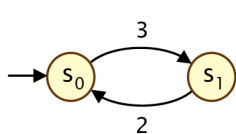


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$\sum_{i=0}^{\infty} e^{-rt} \frac{(rt)^i}{i!} = 1$ due to the fact that $e^{-rt} \frac{(rt)^i}{i!}$ is a (Poisson) distribution

Overview

- 1 Recall: continuous-time Markov chains
- 2 Transient distribution
- 3 Uniformization
- 4 Strong and weak bisimulation
- 5 Computing transient probabilities
- 6 Summary**

Summary

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